Lifetime Risk of Imprisonment Remains High and Starkly Unequal in the United States*

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How likely are Americans to be imprisoned over the course of their lives? Applying life table methods to new survey data from 2016, the first data available since 2004, we calculate the lifetime risk of imprisonment among U.S. males and females of different ethnoracial groups. Black males' cumulative risk of imprisonment remains extremely high but has declined relative to estimates using data from the 1990s and early 2000s. The lifetime risk of imprisonment among American Indians and Alaska Natives—reported here for the first time—is comparable to or higher than that of any other ethnoracial group. Although national prison admission rates are declining, imprisonment remains a pervasive and deeply unequal life-course experience.

Keywords: Race/ethnicity; incarceration; demography; inequality

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The rise of mass incarceration in the United States fundamentally transformed the life course of young men, especially those who were Black, did not finish high school, or lived in poor neighborhoods. This transformation made what was once an extremely rare event—prison incarceration—a common experience [1, 2, 3, 4]. As incarceration became increasingly common, so too did the incarceration of family members [5, 6]. According to recent estimates, just under one in two American women has ever had a family member incarcerated, with risks approaching two in three for some groups [7]. Prisons and jails have come to pervade the lives of the most disadvantaged and dispossessed Americans.

However, existing estimates of the proportion of the American population that will ever experience imprisonment are limited in two important ways. First, they are based on data from 2004 or earlier. As a result, the most current estimates of the cumulative risk of imprisonment fail to capture recent declines in the total U.S. imprisonment rate [8], as well changes in inequality in prison admissions [9, 10]. Second, prior estimates exist only for single-race non-Hispanic White, single-race non-Hispanic Black, and Hispanic people. No previous work has estimated the lifetime risk of imprisonment among American Indians and Alaska Natives (AIAN) or Asian Americans and Pacific Islanders (AAPI). In this article, we use synthetic cohort life tables [11] and restricted-access data from the 2016 Survey of Prison Inmates¹ to present new estimates of the lifetime risk of imprisonment among a more expansive set of ethnoracial groups.

We measure the cumulative risk of imprisonment among male and females of all races and ethnicities and separately among seven non-exclusive and non-exhaustive ethnoracial groups. Following prior research [12], we construct AAPI, Black, and White groups that include only single-race non-Hispanic people, as well as a Hispanic group that includes people of all racial identities. Because a large proportion of AIAN people identify as multiracial and/or Hispanic, we report estimates for three AIAN groups: single-race non-Hispanic AIAN; multiracial and/or Hispanic AIAN; and all people identifying as AIAN, irrespective of other

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racial identities or Hispanic ethnicity. Figure 1 shows that with the possible exception of single-race non-Hispanic AAPI and single-race non-Hispanic AIAN people, there are enough people in the survey to generate stable group estimates. Table 1 illustrates the synthetic cohort life table parameters used to generate estimates of age-specific cumulative risk of imprisonment.²

[Fig. 1 about here.]

[Table 1 about here.]

Fig. 2 reports estimates of the lifetime risk of imprisonment in the United States in 2016, by sex and ethnoracial group. The lifetime risk of imprisonment for all U.S. residents was 6.1%. One in 10 males (10.0%) and 1 in 45 females (2.2%) were imprisoned over the course of their lives. The cumulative risk of imprisonment continued to meaningfully increase well beyond the ages at which rates of prison admission are highest. For example, across all groups, the cumulative risk of imprisonment was 3.7% at ages 30–34, compared to 5.1% at ages 40–44 and 5.7% at ages 50–54.

The lifetime risk of imprisonment in the United States is deeply divided. Most strikingly, Fig. 2 shows for the first time that AIAN people's lifetime risk of imprisonment was similar to or greater than risk of imprisonment among any other ethnoracial group. Considering people only identifying as AIAN, the male imprisonment risk (14.4%) was comparable to the risk among Hispanic males (14.3%) and Black males (17.6%), and the female risk (5.0%) was more than twice as high as that of any other female group. When all people identifying as AIAN are considered, the risk was much higher: nearly half of all AIAN males (48.3%) were imprisoned over the course of their lives, as were nearly 1 in 7 of all AIAN females (15.0%).

²Materials and methods are available as supporting material.

Lifetime prevalence of imprisonment among multiracial and/or Hispanic AIAN people was particularly extreme, reaching 66.0% among males and 21.6% among females.

Fig. 2 is consistent with a large body of prior evidence highlighting Black males' high likelihood of imprisonment. However our estimates suggest that Black males likelihood of imprisonment is falling. Using similarly sources and methods, Bonczar and Beck [1] and Bonzcar [2] report estimates of Black men's lifetime risk of imprisonment between 28.5% and 32.2%, compared to our estimate of 17.6%. Using slightly different methods, Pettit and Western [3] estimate Black men's cumulative risk of imprisonment up to ages 30 to 34. They estimate risks between 18.7% and 20.5%, compared to our estimate of 12.1%. Still, despite these declines, in 2016 Black men's likelihood of being imprisoned over the life course was still greater than 1 in 8.

A comparable percentage of Hispanic males ever experienced imprisonment (14.3%). While much lower, the lifetime risk of imprisonment among White males (6.2%) was higher than estimated using data from the 1990s and early 2000s [1, 2, 3], and among AAPI males (2.7%) was nontrivial. Our estimates of the cumulative risk of imprisonment among females reveal previously undocumented patterns of inequality. Black females' cumulative risk of imprisonment exceeded that of White females early in the life course, but the risk for White females overtook the risk for Black females by ages 50-54. The lifetime risk of imprisonment among all AIAN females was more than 7 times greater than among the next most frequently imprisoned group, White females, a disparity much larger than that between the corresponding male groups (2.7:1). Measured using a relative risk, sexual inequality in the lifetime risk of imprisonment was greatest among Black people (8.8:1).

Our results support several conclusions. First, the cumulative risks of imprisonment we estimate continue to support a characterization of the United States' criminal justice system as mass incarceration [13, 14]. After decades of steady increases, rates of imprisonment in the United States began declining in 2008 [8]. But imprisonment rates remain at historically high levels, and this is reflected in our estimates of Americans' very high lifetime risk of

imprisonment. Prison incarceration is still an extremely common life-course event in the United States, particularly among its most disadvantaged groups.

Second, new evidence that American Indian and Alaska Native people are among the primary targets of contemporary mass incarceration should encourage new thinking about its contours, causes, and consequences. The sources of Black Americans' unequal incarceration have been widely explored [14, 15]. A historical legacy of violence, enslavement, dispossession, segregation, and discrimination has predisposed contemporary criminal justice institutions against equal justice for Black people. This legacy has furthermore created social and economic disadvantages that compound Black people's risk of imprisonment. Though distinctive in important ways, American Indians and Alaska Natives share this historical legacy of subjection to the most severe forms of state-sanctioned social control. Research should consider the extent to which existing explanations for the mass incarceration of Black Americans may or may not extend to American Indians and Alaska Natives.

One clear difference is that American Indian and Alaska Native people are uniquely subject to a complex system of laws, police, courts, and correctional facilities that combines local, state, federal, and tribal jurisdiction. Many criminal jurisdictions in the United States are reducing prosecutions and prison sentences. By contrast, the Tribal Law and Order Act of 2010 increased tribal court sentencing authority and resulted in large increases in federal prosecutions of criminal cases in Indian country. Jurisdictional variation in criminal justice reform has been proposed as an explanation for narrowing Black-White inequality and widening geographic inequality in incarceration [16, 17, 18]. Our estimates of the cumulative risk of imprisonment among AIAN people recommends greater attention to jurisdictional variation in criminal justice reform as a source of contemporary inequality in mass incarceration.

Our finding of marked heterogeneity in imprisonment risk among people identifying as AIAN makes clear that no single estimate neatly summarizes this group's experience of the criminal justice system. Differences in risk of imprisonment between people identifying only as AIAN and AIAN people identifying as multiracial and/or Hispanic may reflect the

differential likelihood of facing tribal or federal jurisdiction. But it may also reflect the well-documented fluidity of reported ethnoracial identification among people identifying as AIAN. American Indian and Alaska Native racial identification is uniquely affected by factors such as tribal governance, indigenous homelands, and tribal enrollment blood quantum requirements. But growing rates of reported AIAN identification in census and survey data far exceed expectations based on mortality and migration patterns, and people reporting partial AIAN race are more likely than people reporting other races and ethnicities to shift their reported identification with AIAN race, other races, and Hispanic ethnicity [19, 20]. Care is needed in research to attend to the diversity of identities and experiences within this group, and how identification itself may be related to criminal justice system contact.

Finally, the extreme risk of imprisonment among American Indian and Alaska Native females documented in our study demands special attention. Rates of violent victimization among AIAN women are comparable to or higher than rates among other ethnoracial groups [21, 22, 23]. Our findings show that AIAN females are also at outlying risk of experiencing prison incarceration. Ensuring safety and justice for American Indian and Alaska Native women without subjecting them to extreme levels of social control should be a priority of criminal justice policy and practice in the United States.

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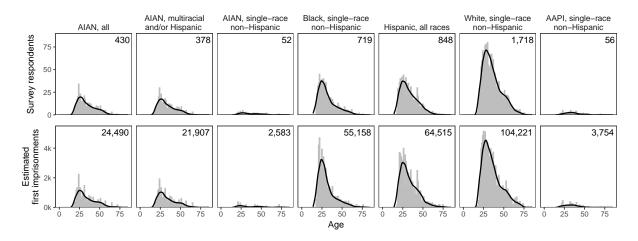


Fig. 1 Number and age distribution of sampled survey respondents (top row) and estimated first imprisonments (bottom row), by ethnoracial group. Respondents are people observed in the 2016 Survey of Prison Inmates who were admitted to prison for the first time within the previous year. Population estimates are calculated using survey weights and an adjustment for the underestimation of prison admissions due to short prison spells. Data limitations prevent separate estimates for multiracial non-Hispanic people. AAPI, Asian American/Pacific Islander; AIAN, American Indian/Alaska Native.

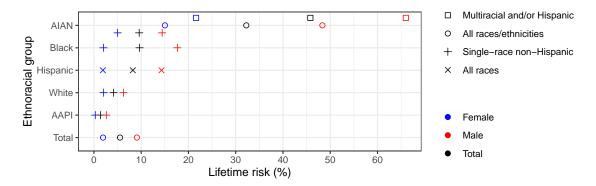


Fig. 2 Lifetime risk of imprisonment, by sex and ethnoracial group. Total estimates on the bottom row include people of all ethnoracial groups; total estimates plotted as circles include people of both sexes. AAPI, Asian American/Pacific Islander; AIAN, American Indian/Alaska Native.

Table 1: Synthetic cohort life table, all males and females identifying as American Indian or Alaska Native, United States, 2016.

Age interval	$_{n}N_{x}$	$_{n}N_{x}^{\prime}$	$_{n}\hat{D}_{x}$	$_nm_x$	$_{n}a_{x}$	$_{n}q_{x}$	$_{n}p_{x}$	l_x	$_nd_x$	c_x
15-19	454,643	454,643	390	0.0009	2.5	0.0043	0.9957	100,000	429	0.0043
20-24	435,795	433,927	6,334	0.0146	2.5	0.0704	0.9296	99,571	7011	0.0744
25-29	384,781	356,266	4,212	0.0118	2.5	0.0574	0.9426	$92,\!560$	5315	0.1275
30-34	372,707	326,702	3,886	0.0119	2.5	0.0578	0.9422	87,246	5039	0.1779
35-39	$335,\!564$	279,198	3,003	0.0108	2.5	0.0524	0.9476	82,206	4305	0.2210
40-44	341,028	271,326	1,710	0.0063	2.5	0.0310	0.9690	77,901	2417	0.2452
45-49	$337,\!549$	261,908	1,998	0.0076	2.5	0.0374	0.9626	$75,\!484$	2825	0.2734
50-54	361,737	$272,\!492$	$1,\!453$	0.0053	2.5	0.0263	0.9737	72,660	1911	0.2925
55-59	318,871	$235,\!426$	904	0.0038	2.5	0.0190	0.9810	70,748	1346	0.3060
60-64	261,863	190,615	107	0.0006	2.5	0.0028	0.9972	69,403	195	0.3079
65-69	190,468	138,362	316	0.0023	2.5	0.0113	0.9887	69,208	786	0.3158
70-74	126,893	91,418	177	0.0019	2.5	0.0096	0.9904	68,422	658	0.3224
75-79	78,265	55,993	0	0.0000	2.5	0.0000	1.0000	67,765	0	0.3224
80-84	50,875	36,398	0	0.0000	2.5	0.0000	1.0000	67,765	0	0.3224
85+	41,048	29,367	0	0.0000	2.5	0.0000	1.0000	67,765	0	0.3224

Notes: ${}_{n}N_{x}$ is the number of U.S. residents in the age interval x to x+n; ${}_{n}N'_{x}$ is the number of people who were at risk of first imprisonment in the interval; ${}_{n}\hat{D}_{x}$ is the number people who experienced first imprisonment in the interval; ${}_{n}m_{x}$ is the rate at which people experienced first imprisonment in the interval; ${}_{n}a_{x}$ is the assumed number of years in the interval to first imprisonment among those who experienced first imprisonment in the interval; ${}_{n}q_{x}$ is the probability of imprisonment in the interval conditional on not having been imprisoned before the interval; ${}_{n}p_{x}$ is the probability of not being imprisoned in the interval conditional on not having been previously imprisoned; ${}_{l}x$ is the number of people in a hypothetical cohort of 100,000 people who were not imprisoned by the beginning of the interval; ${}_{n}d_{x}$ is the number of people in the hypothetical cohort who experienced first imprisonment in the interval; and ${}_{l}x$ is the cumulative probability of imprisonment by the end of the interval.

Supporting Material: Lifetime Risk of Imprisonment Remains High and Starkly Unequal in the United States*

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Measurement of annual first prison admissions

Calculating the cumulative risk of imprisonment using synthetic cohort life table methods [1] requires age-specific estimates of the annual number of people imprisoned for the first time. To estimate this number we use restricted-access data from the 2016 Survey of Prison Inmates (SPI). The SPI is a nationally representative survey of inmates in state and federal correctional facilities. To identify first imprisonments, we exclude respondents who (i) reported more than one lifetime admission to prison; (ii) reported a first prison admission date earlier than the reported admission date for their current prison spell; or (iii) were comitted for a parole violation.

To identify annual first imprisonments, we further exclude respondents admitted to prison more than one year before being interviewed. To protect imprisoned people's identities, the

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Bureau of Justice Statistics (BJS) suppresses interview dates, even in restricted-access files. Therefore, the main estimates rely on an indirect method of estimating interview date—assigning the first day of the known weighted mean interview month (June 2016) to all records. Mean interview month is calculated using tabulations provided directly by BJS. Two pieces of evidence suggest that this method does not significantly bias our estimates.

First, the main estimates for 2016 are very similar to estimates based on an alternative interview date estimation strategy—assuming that all the interviews in a given correctional facility took place on the first day of the month following the latest admission month recorded among all survey respondents in that facility. Figure A1 compares the main and alternative approaches, plotting ratios of age-specific estimates of cumulative prevalence of imprisonment. The main approach yields estimates that are broadly conservative compared to the alternative, facility-based approach. Age-specific estimates are higher among people aged 15–19, among whom prevalence is very low. Estimates of lifetime prevalence are lower for all groups except single-race non-Hispanic AIAN males and single-race non-Hispanic White females.

Second, estimates of cumulative prevalence of imprisonment are very similar when using our mean month interview estimation strategy and actual interview dates in another survey wave. Figure A2 uses data from the 2004 Surveys of Inmates of State and Federal Corrections Agencies, plotting ratios of estimates using estimated and actual interview dates for 2004. The largest ratio in lifetime estimates, for single-race non-Hispanic Asian American and Pacific Islander females, was 1.031:1.

Using standard interval notation, for each sex and ethnoracial group we estimate the national number of annual first prison admissions between ages x and x + n as:

$$_{n}\hat{D}_{x} = \sum_{i} (w_{i} \times \gamma),$$

where w_i is an individual survey weight calculated by the Bureau of Justice Statistics (BJS) and γ is an adjustment factor that corrects for the undercounting of annual prison admissions.

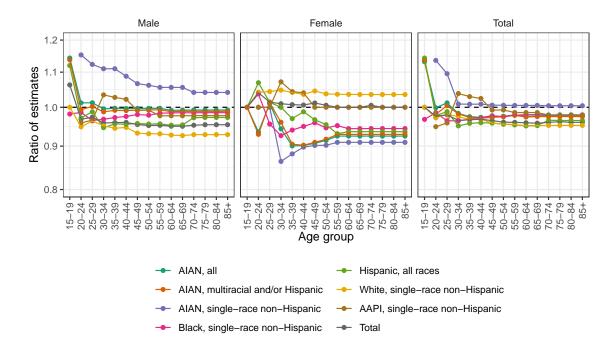


Fig. A1 Ratio of main estimates of age-specific cumulative prevalence of imprisonment to estimates based on an alternative method for estimating interview dates. The main method assigns the first day of the weighted mean interview month to all respondents. The alternative method assigns all respondents in a given facility the first day of the month following the latest recorded admission month in that facility. The vertical axis is plotted on a logarithmic scale.

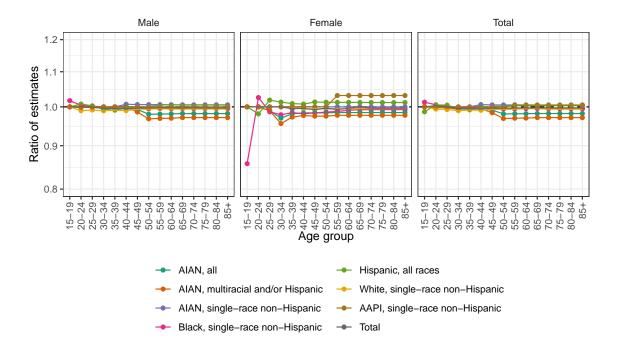


Fig. A2 Ratio of estimates of age-specific cumulative prevalence of imprisonment in 2004 using the main interview estimation strategy (weighted mean month) and actual interview dates. The vertical axis is plotted on a logarithmic scale.

Among people admitted to prison in the year before interviews were conducted, those who had already been released were not interviewed. Therefore, following Bonzcar and Beck [2], Bonzcar [3], and Pettit and Western [4], we use restricted-access data from the National Corrections Reporting Program (NCRP)¹ to calculate adjustment factors to account for this negative bias. Similarly to prior studies, we restrict calculations to new court commitments with sentences greater than one year. We use the NCRP term files, which provide the most reliable measures of admissions and releases. The term files include records from all states except Arkansas, Connecticut, Hawaii, Idaho, Maryland, North Dakota, Oregon, Vermont, Virginia, and West Virginia.

For each survey year t and jurisdiction j (state or federal), we calculate adjustment factors such that:

$$\gamma_{t,j} = \frac{admissions_{t,j}}{admissions_{t,j} - releases_{t,j}},$$

where $admissions_{t,j}$ is the number of people admitted to prison in the previous twelve months and $releases_{t,j}$ is the number of these people released within twelve months of admission. NCRP data lack complete information on federal admissions and releases after 1991, but state and federal prisons have different rates of turnover. To account for this, we multiply annual adjustment factors for 2016 calculated only using state data by the ratio of factors for the 1990 calendar year calculated using, respectively, state and federal data combined and state data only:

$$\gamma = \gamma_{2016,state} \times \frac{\gamma_{1990,combined}}{\gamma_{1990,state}}$$
.

Figure 1 in the main text reports the number and age distribution of sampled survey respondents and estimated first imprisonments ethnoracial group. Figure A3 reports the same distributions further disaggregated by sex.

¹Available at (https://www.icpsr.umich.edu/web/NACJD/series/38).

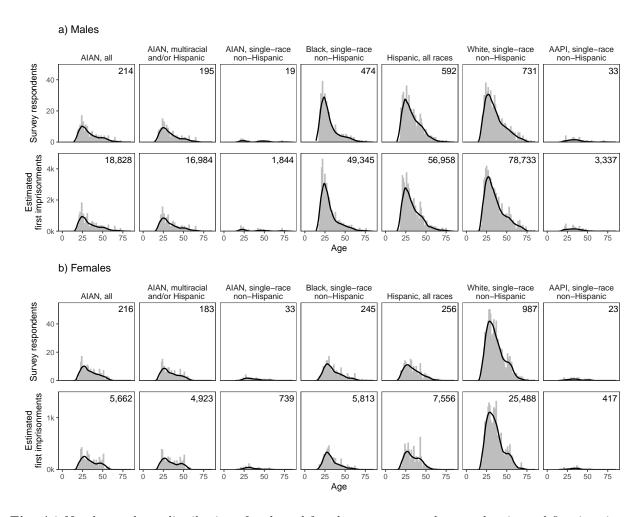


Fig. A3 Number and age distribution of male and female survey respondents and estimated first imprisonments, by ethnoracial group. Respondents are people observed in the 2016 Survey of Prison Inmates who were admitted to prison for the first time within the previous year. Population estimates are calculated using survey weights and an adjustment for the underestimation of prison admissions due to short prison spells, calculated using data from the National Corrections Reporting Program. AAPI, Asian American/Pacific Islander; AIAN, American Indian/Alaska Native.

Calculation of life tables

To calculate the cumulative risk of imprisonment over the life course, we use synthetic cohort life table methods [1]. We calculate life tables separately by sex, by ethnoracial group, and by sex and ethnoracial gruop. We use five-year age intervals beginning at ages 15–19, except for ages 85 and older, which we combine into a single group. ${}_{n}N_{x}$ is the mid-year U.S. residential population aged x to x + n, measuring using 2016 American Community Survey (ACS) five-year samples.² Because the ACS includes people residing in institutions, our estimates of the cumulative risk of imprisonment will be conservative. $_{n}\hat{D}_{x}$ is the number of first imprisonments in the age interval, estimated as described above. $_{n}a_{x}$ is, among those imprisoned for the first time during the interval, the average number of years in the interval until imprisonment, assumed to be the midpoint (2.5). First, we preliminarily estimate several standard life table parameters. $_{n}m_{x}$ is the rate of first imprisonment in the interval; $_{n}q_{x}$ is the probability of first imprisonment in the interval contingent upon not being imprisoned before the interval; and $_{n}p_{x}$ is the probability of not experiencing first imprisonment contingent on not being imprisoned before the interval, calculated respectively as:

$$_{n}m_{x} = \frac{_{n}\hat{D}_{x}}{_{n}N_{x}};\tag{1}$$

$$_{n}q_{x} = \frac{n \cdot {}_{n}m_{x}}{1 + (n - {}_{n}a_{x}) \cdot {}_{n}m_{x}}, \quad {}_{\infty}q_{85} = 1;$$
 (2)

$$_{n}p_{x}=1-_{n}q_{x}. \tag{3}$$

 $_{n}N_{x}$, the U.S. residential population, includes people who have already experienced imprisonment, yielding downwardly biased estimates of life table parameters. To calculate the age-specific population at risk of imprisonment, we multiply $_{n}N_{x}$ by the likelihood of not being imprisoned before age x:

$${}_{n}N'_{x} = \begin{cases} {}_{n}N_{x}, & x = 15 \\ {}_{n}N_{x} \cdot \prod_{j=15}^{x-n} {}_{n}p_{j}, & x > 15. \end{cases}$$

Then, we substitute ${}_{n}N'_{x}$ for ${}_{n}N_{x}$ in Equation 1 and recalculate Equations 1 through 3. We then define l_{x} as a hypothetical cohort of 100,000 people who will not have been imprisoned by the beginning of the interval, and ${}_{n}d_{x}$ as the number of people in the hypothetical cohort

²Available at (https://usa.ipums.org/usa/).

Table A1: Cumulative risk of imprisonment (%) by age and ethnoracial group, males and females combined, United States, 2016.

Age group	Total	AIAN,	AIAN, multiracial and/or Hisp.	AIAN, single-race non-Hisp.	Black, single-race non-Hisp.	Hispanic, all races	White, single-race non-Hisp.	AAPI, single-race non-Hisp.
15-19	0.09	0.43	0.67	0.00	0.30	0.16	0.03	0.00
20-24	1.22	7.44	10.45	2.51	2.69	1.75	0.69	0.37
25-29	2.41	12.75	18.63	3.33	5.10	3.36	1.57	0.46
30-34	3.32	17.79	25.84	4.56	6.67	4.44	2.33	0.82
35-39	4.05	22.10	32.47	4.84	7.62	5.39	3.00	1.05
40-44	4.56	24.52	35.74	5.84	8.15	6.31	3.40	1.27
45-49	4.88	27.34	39.57	6.99	8.64	6.87	3.64	1.33
50-54	5.10	29.25	42.39	7.39	8.93	7.28	3.83	1.36
55-59	5.27	30.60	44.21	7.94	9.20	7.55	3.98	1.39
60-64	5.36	30.79	44.52	7.94	9.37	7.76	4.04	1.39
65-69	5.44	31.58	45.79	7.94	9.48	7.88	4.12	1.39
70-74	5.50	32.24	45.79	9.58	9.52	8.08	4.16	1.41
75-79	5.53	32.24	45.79	9.58	9.65	8.23	4.16	1.41
80-84	5.53	32.24	45.79	9.58	9.65	8.23	4.17	1.41
85+	5.53	32.24	45.79	9.58	9.65	8.23	4.17	1.41

Notes: AAPI, Asian American/Pacific Islander; AIAN, American Indian/Alaska Native; Hisp., Hispanic.

first imprisoned in the interval:

$$l_0 = 100,000, \quad l_{x+n} = l_x \cdot {}_{n}p_x;$$

 ${}_{n}d_x = l_x - l_{x+n}..$

Cumulative risk of imprisonment ${}_{n}c_{x}$ can then be calculated as:

$$_{n}c_{x} = \frac{\sum_{j=15}^{x} {_{n}d_{j}}}{l_{0}}.$$

Tables A1, A2, and A3 report age-specific estimates of cumulative risk of imprisonment for both sexes combined, for males, and for females, respectively. Estimates of $_{\infty}c_{85}$ in the bottom row of each table correspond to the estimates of lifetime risk reported in Figure 2 of the main text.

Table A2: Cumulative risk of imprisonment (%) by age and ethnoracial group, males, United States, 2016.

Age group	Total	AIAN,	AIAN, multiracial and/or Hisp.	AIAN, single-race non-Hisp.	Black, single-race non-Hisp.	Hispanic, all races	White, single-race non-Hisp.	AAPI, single-race non-Hisp.
15-19	0.17	0.80	1.24	0.00	0.54	0.28	0.04	0.00
20-24	2.12	11.75	16.01	4.38	5.01	3.06	1.10	0.69
25-29	4.03	19.81	28.47	5.19	9.29	5.72	2.38	0.85
30-34	5.50	26.72	38.64	5.84	12.13	7.56	3.51	1.51
35-39	6.61	33.27	48.46	5.84	13.77	9.13	4.43	1.89
40-44	7.40	36.40	52.40	7.18	14.73	10.59	5.00	2.35
45-49	7.93	40.58	57.50	9.34	15.59	11.57	5.37	2.47
50-54	8.28	42.96	60.63	10.02	16.14	12.27	5.64	2.52
55-59	8.58	45.01	63.15	10.86	16.68	12.80	5.89	2.57
60-64	8.74	45.39	63.70	10.86	16.98	13.22	6.00	2.57
65-69	8.89	46.97	66.00	10.86	17.23	13.48	6.14	2.57
70-74	9.02	48.34	65.99	14.44	17.33	13.94	6.23	2.62
75-79	9.08	48.34	65.99	14.44	17.65	14.30	6.23	2.62
80-84	9.09	48.34	65.99	14.44	17.65	14.30	6.24	2.62
85+	9.09	48.34	65.99	14.44	17.65	14.30	6.24	2.62

Notes: AAPI, Asian American/Pacific Islander; AIAN, American Indian/Alaska Native; Hisp., Hispanic.

Table A3: Cumulative risk of imprisonment (%) by age and ethnoracial group, females, United States, 2016.

Age group	Total	AIAN, all	AIAN, multiracial and/or Hisp.	AIAN, single-race non-Hisp.	Black, single-race non-Hisp.	Hispanic,	White, single-race non-Hisp.	AAPI, single-race non-Hisp.
15-19	0.02	0.03	0.05	0.00	0.05	0.03	0.01	0.00
20-24	0.26	2.60	3.94	0.54	0.28	0.31	0.25	0.04
25-29	0.71	5.03	7.44	1.38	0.86	0.75	0.72	0.07
30-34	1.05	8.13	11.34	3.16	1.25	0.98	1.10	0.15
35-39	1.41	10.04	14.12	3.71	1.57	1.28	1.52	0.24
40-44	1.62	11.69	16.43	4.38	1.69	1.63	1.74	0.26
45-49	1.74	13.08	18.63	4.58	1.86	1.75	1.85	0.27
50-54	1.84	14.45	20.85	4.70	1.92	1.86	1.96	0.29
55-59	1.89	15.02	21.62	5.01	1.96	1.88	2.02	0.29
60-64	1.91	15.02	21.62	5.01	2.01	1.90	2.04	0.29
65-69	1.92	15.02	21.62	5.01	2.01	1.90	2.05	0.29
70-74	1.93	15.02	21.62	5.01	2.02	1.90	2.06	0.29
75-79	1.93	15.02	21.62	5.01	2.02	1.90	2.06	0.29
80-84	1.93	15.02	21.62	5.01	2.02	1.90	2.06	0.29
85+	1.93	15.02	21.62	5.01	2.02	1.90	2.06	0.29

Notes: AAPI, Asian American/Pacific Islander; AIAN, American Indian/Alaska Native; Hisp., Hispanic.

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